

Poisson

Geometric

Memoryless property.

4.5 Exponential distribution

Recall $X \sim \text{Geom}(p) \rightarrow \{1, 2, \dots\}$ (discrete random variable)
 = "waiting time until first success".

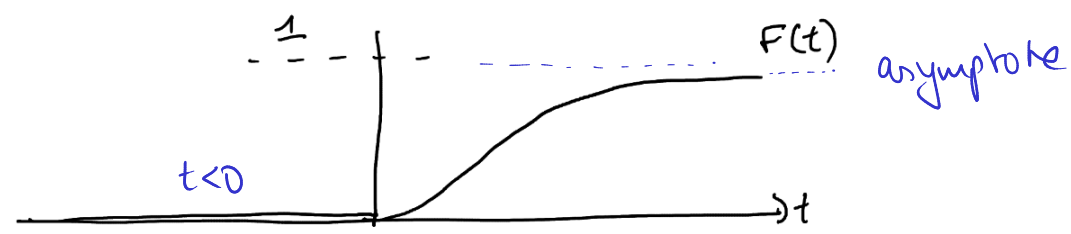
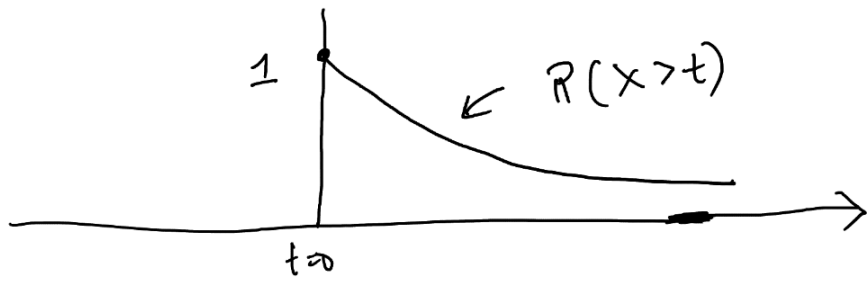
But $X = 1, 2, 3, \dots$ etc so it is discrete.

$X \sim \text{Exp}(\lambda)$ is a waiting time that is continuous.

cdf
 $F(t) = P(X \leq t)$

$$P(X > t) = e^{-\lambda t} \quad (\text{Tail}) \quad t \geq 0. \quad \lambda \text{ is some parameter like } 4.$$

$$1 - F(t) \iff F(t) = 1 - e^{-\lambda t} \quad t \geq 0.$$



The pdf is the derivative of the cdf.

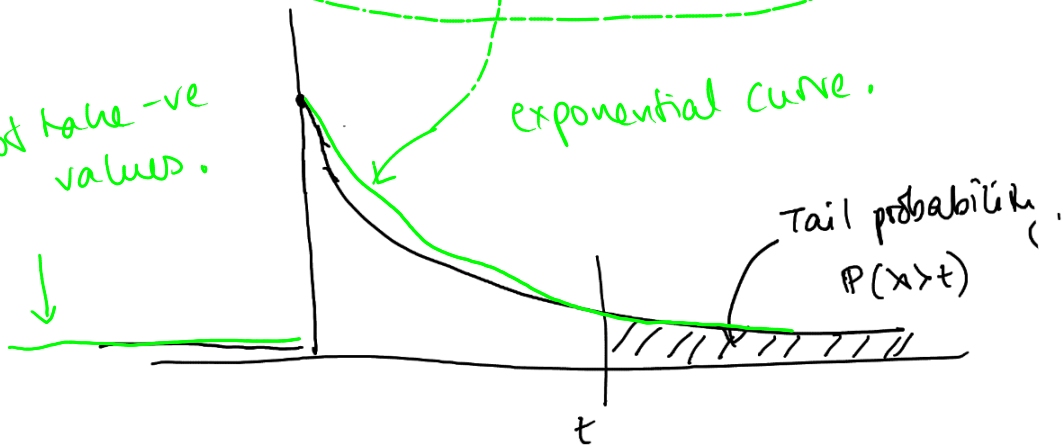
$$F(t) = 1 - e^{-\lambda t} \quad t \geq 0$$

Therefore: $f(t) = \frac{d}{dt} F(t)$. (use the chain rule)

$$f(t) = \begin{cases} \lambda e^{-\lambda t} & t \geq 0 \\ 0 & t < 0 \end{cases}$$

write down in your cheatsheet

X cannot take -ve values.



X models a waiting time $\Rightarrow X \geq 0$

Expectation and variance

$$E[X] = \int_0^{\infty} \underbrace{t}_{\text{(value)}} \underbrace{\lambda e^{-\lambda t}}_{\text{(probability)}} dt$$

differentiate this away

(When in doubt)
↳ you integrate by parts.

$$u = t \quad dv = \lambda e^{-\lambda t} dt \Rightarrow v = (-e^{-\lambda t}) \quad \text{pdf.}$$

$$= t \left(\frac{0}{-e^{-\lambda t}} \right) \Big|_0^{\infty} - \int_0^{\infty} 1 \cdot (-e^{-\lambda t}) dt = \int_0^{\infty} e^{-\lambda t} dt = \frac{e^{-\lambda t}}{-\lambda} \Big|_0^{\infty}$$

$$= \frac{1}{\lambda}$$

$$\text{Var}(X) = \frac{1}{\lambda^2}$$

↓

(we will compute this more efficiently using mgfs later)

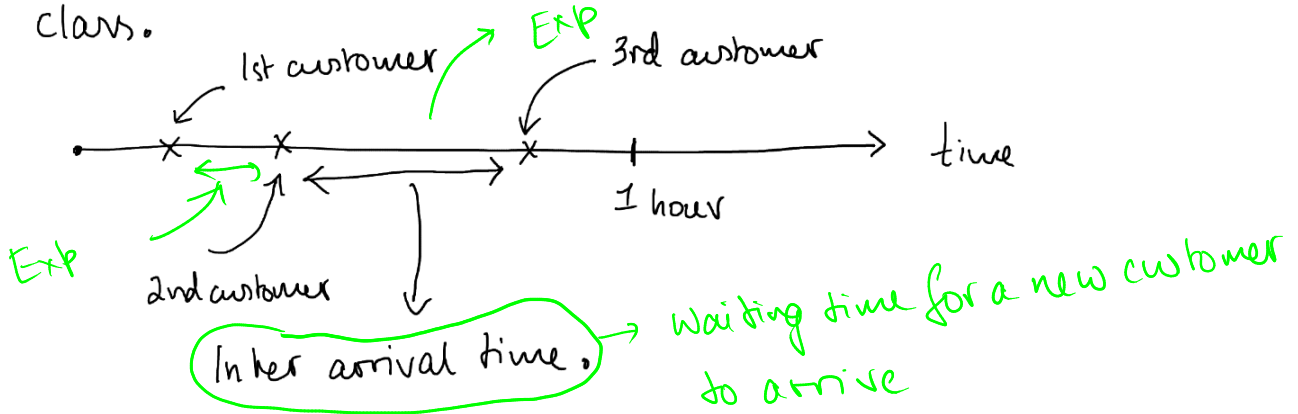
You have to integrate by parts

twice.

Apply HW.

Exponential and Poisson

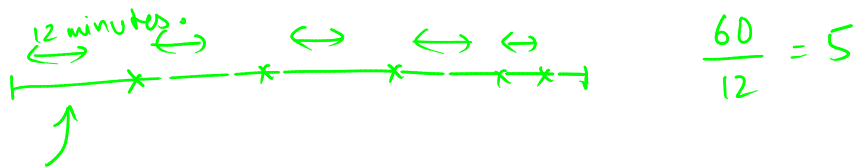
We will not cover this in detail in this class.



This interarrival time is $\text{Exp}(\lambda)$.

The inter arrival times are also independent.

If $\lambda = 5$ customers/hour. Then the average time between customers coming in should be $\frac{1}{5}$ hour, or every 12 minutes.



$$Y \sim \text{Exp}(\lambda) \quad E[Y] = \frac{1}{\lambda} = \frac{1}{5} \quad (\text{as required})$$

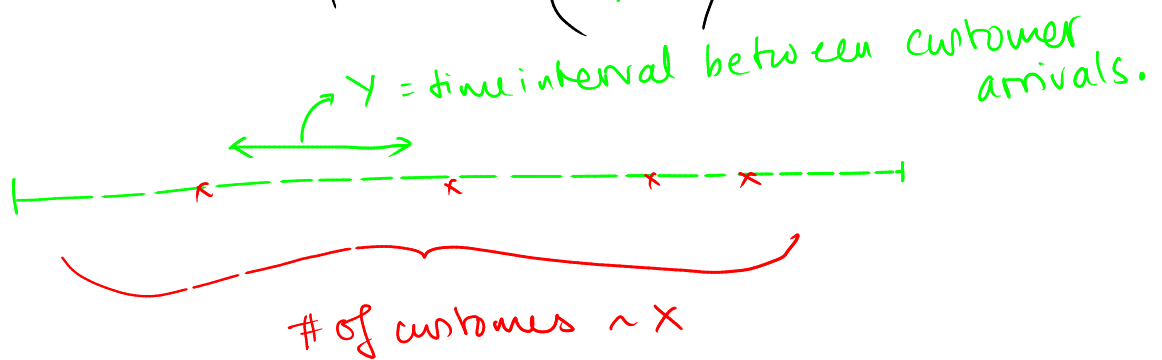
Let

$X =$ # of customers that come to the store
on one day (say Jan 7)

Assume $X \sim \text{Poisson}(\lambda)$

Let $Y =$ Time between successive customer arrivals.

Then $Y \sim \text{Exponential}(\lambda)$



Memory less property

$X \sim$ waiting time.
 $\sim \text{Exp}(\lambda)$

Suppose you have waited 3 minutes already. What is the probability that you wait 5 more minutes?

Let's unpack that. $X =$ waiting time.

waited 3 minutes and it has not

already happened : $X > 3$ ^{given}

Probability of having to wait at least 5 more minutes : $X > 5 + 3$.

This is what you're asked to compute.

$$\begin{aligned} & \leftarrow P(X > 5 + 3 \mid X > 3) \\ & = \frac{P(X > 5 + 3 \cap X > 3)}{P(X > 3)} \end{aligned}$$

$$P(A \mid B) = \frac{P(A \cap B)}{P(B)}$$

$$= \frac{P(X > 8)}{P(X > 3)}$$



=

$$P(X > t) = e^{-\lambda t}$$

$$P(X > 5 + 3 \mid X > 3) = \frac{e^{-\lambda 8}}{e^{-\lambda 3}} = e^{-5\lambda} = P(X > 5)$$

Geometric and Exponential

Suppose

$$X \sim \text{Exp}(\lambda)$$

Then $P(X > t) = e^{-\lambda t}$

Suppose

$$Y \sim \text{Geometric}(p)$$

Then $P(Y > t) = P(\text{1st } t \text{ trials are failures})$

$$= (1-p)(1-p) \dots (1-p)$$

$$= (1-p)^t = e^{\underbrace{t \log(1-p)}_{t^{\text{th}} \text{ trial}}}$$

$$\lambda = -\log(1-p)$$

Geometrics appropriately scaled \Rightarrow Exponential.

Recall p = probability of success.

INTERESTING | UNSURPRISING FACT:

The exponential is a limit of a geometric random variable!

I won't make this precise, but if



(squeeze points together)

.....
 $\frac{1}{100}, \frac{2}{100}$

then this goes to the continuous random variable $X \sim \text{Exp}(\lambda)$.

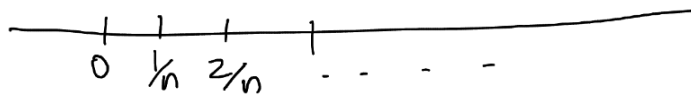
See theorem 4.33 in your text book for details

In case you're really curious, here is a short proof of this fact.

The Geometric distribution converges to exponential

I usually skip this: SAFELY IGNORE.

$$\frac{1}{n} \text{Geom}(p) = \frac{Y_n}{n}$$



$$P\left(\frac{\text{Geom}(p)}{n} > x\right) = \sum_{k=\lceil nx \rceil}^{\infty} P(Y_n = k)$$

$$= \sum_{k=\lceil nx \rceil}^{\infty} (1-p)^{k-1} p$$

$$= (1-p)^{\lceil nx \rceil - 1} \frac{1}{1-p}$$

$$\text{let } p = \frac{\lambda}{n}$$

using the geometric formula.

$$= \left(1 - \frac{\lambda}{n}\right)^{\lceil nx \rceil - 1} \approx \frac{\left(1 - \frac{\lambda}{n}\right)^{nx}}{\left(1 - \frac{\lambda}{n}\right)} = \frac{\left(\left(1 - \frac{\lambda}{n}\right)^n\right)^x}{1 - \frac{\lambda}{n}}$$

$$\rightarrow \frac{(e^{-\lambda})^x}{1} = P(\text{Exp}(\lambda) > x)$$